

# Equal Representation in Two-tier Voting Systems

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# Introduction

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- History and efficiency considerations often call for *two-tier electoral systems*:
  1. People's preferences are aggregated in *constituencies*
  2. Constituencies' preferences are aggregated in an *electoral college*
- Question:

How should constituencies' voting weights in the college be chosen s.t. a priori all individuals have identical influence?
- Allocating weights proportional to population sizes seems straightforward, but:
- In general, voting *power* is not linear in voting *weight*, e.g. EU Council of Ministers 1958.
- Power measures as the Penrose-Banzhaf- or the Shapley-Shubik-Index are designed to capture the non-trivial relationship between weight and power.

## Penrose's square root rule

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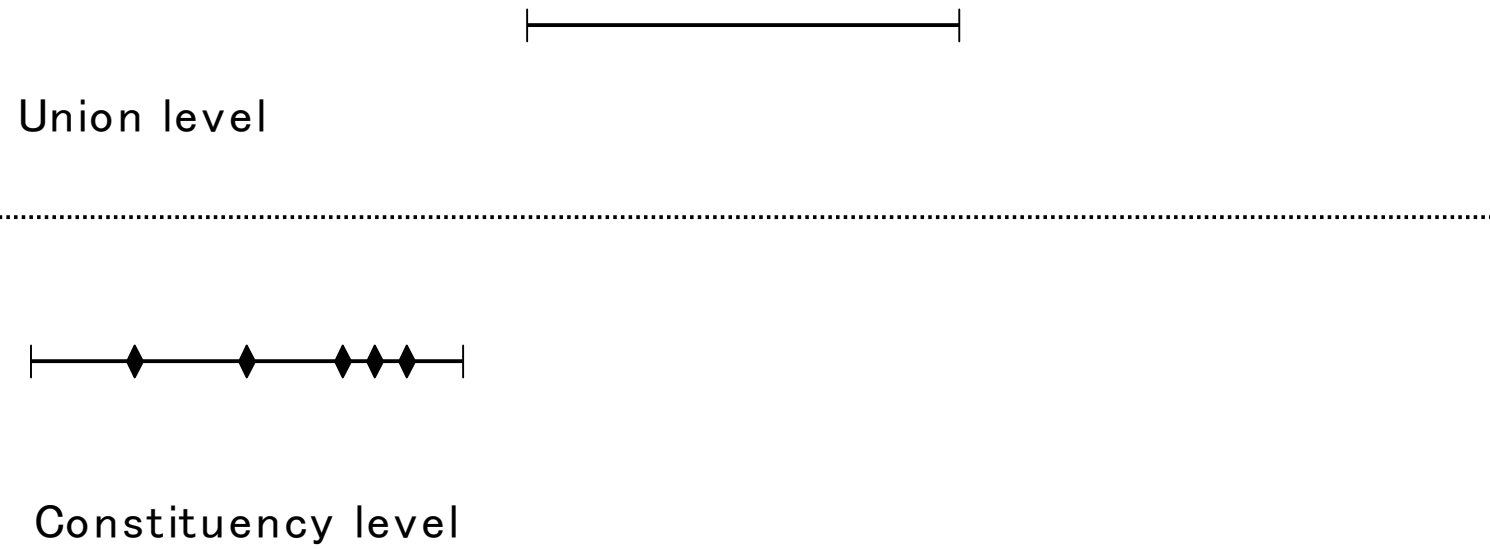
- *Penrose's square root rule (1946):*  
Choose weights s.t. constituencies' Penrose-Banzhaf index is proportional to square root of population
- For most practical reasons (especially, if the number of constituencies is “large”), a simpler rule suffices:  
$$\text{weight} = \text{sqrt}(\text{Population})$$
- The rule requires decisions  $x \in \{0, 1\}$  and (in expectation) equi-probable independent 0 or 1-votes
- What if the world is *not* dichotomous but, e.g.,  $x \in [0, 1]$ ?

# Outline

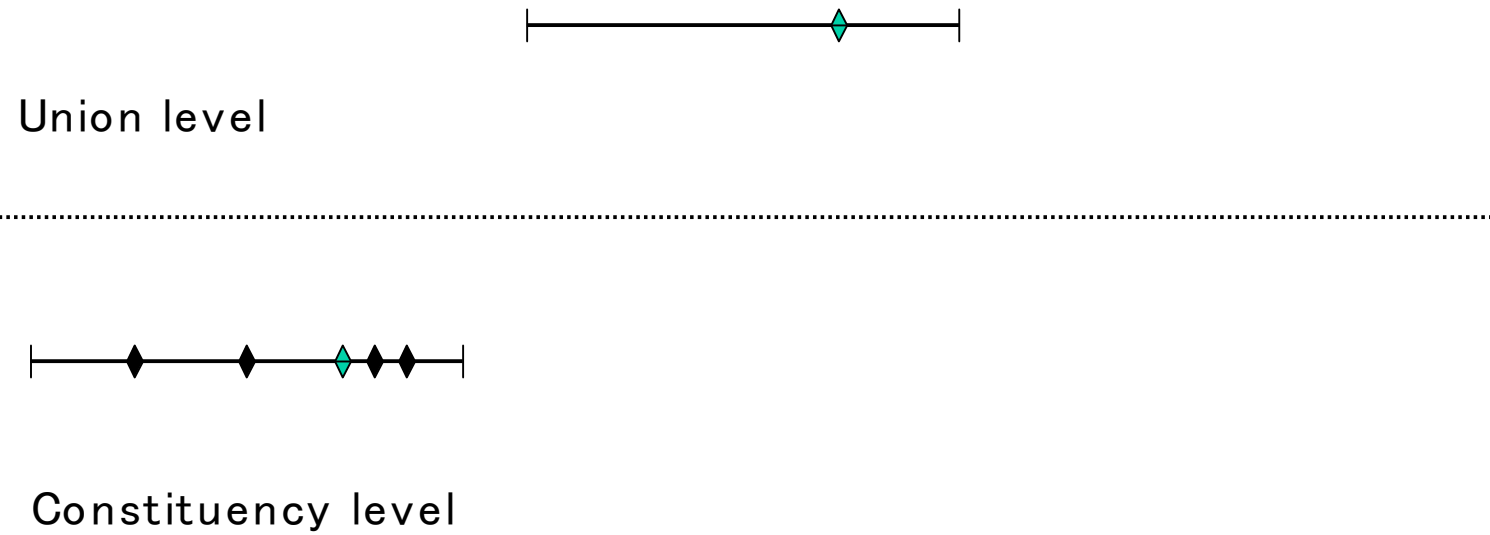
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- Model
- Analytical problems
- Monte Carlo simulation
- Results
- Concluding remarks

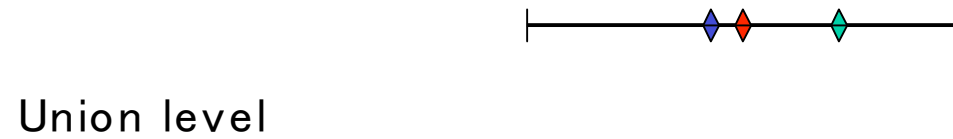
# Model



# Model



# Model



Constituency level

# Model

- Voters are partitioned into  $m$  constituencies and have single-peaked preferences with a priori uniformly distributed ideal points  $\lambda \in X \equiv [0,1]$
  - Constituency  $j$ 's representative is chosen to match the *median voter* in his constituency
  - Each constituency  $j$  has weight  $w_j$  in the electoral college; a 50%-quota  $q$  is used
  - *Pivotal constituency* ( $P$ ) is defined by  $P \equiv \min\{r : \sum_{k=1}^r w_{(k)} > q\}$  [permutation  $(\cdot)$  orders constituencies from left to right]
  - ( $P$ ) gets its will, i.e.  $x^* = \lambda_{(P)}$
- *Problem of equal representation:*  
Given population sizes  $n_1, \dots, n_m$ , find weights  $w_1, \dots, w_n$  s.t. each voter has equal chance of determining  $x^*$



## First analysis

- Each voter in constituency  $j$  has chance  $1/n_j$  to be its median  
 $\Rightarrow \Pr(\lambda_j = \lambda_{P:m}) \leq c \cdot n_j$  for all  $j$  (with  $c > 0$ )
- Assuming *i.i.d.* voters, different  $n_j$  imply different a priori distributions of medians
- With density  $f$  and c.d.f.  $F$  for individual voters' ideal points, representatives' ideal points are asymptotically normal with

$$\mu_j = F^{-1}(0.5), \quad \sigma_j = [2f(\mu_j) \cdot \text{sqrt}(n_j)]^{-1}$$

$\Rightarrow$  Larger constituencies are a priori more central in the electoral college and more likely to be pivotal under a 50%-quota.

## Analytical problems

- Already for *unweighted* voting, i.e.  $P \equiv (m+1)/2$ , we run into trouble:

$$\begin{aligned}\Pr(j = (p)) &= \Pr(\text{exactly } p-1 \text{ of the } \lambda_k, k \neq j, \text{ satisfy } \lambda_k < \lambda_j) \\ &= \int \sum_{\substack{S \subset N \setminus j, \\ |S|=p-1}} \prod_{k \in S} F_k(x) \cdot \prod_{k \in N \setminus j \setminus S} (1 - F_k(x)) \cdot f_j(x) dx\end{aligned}$$

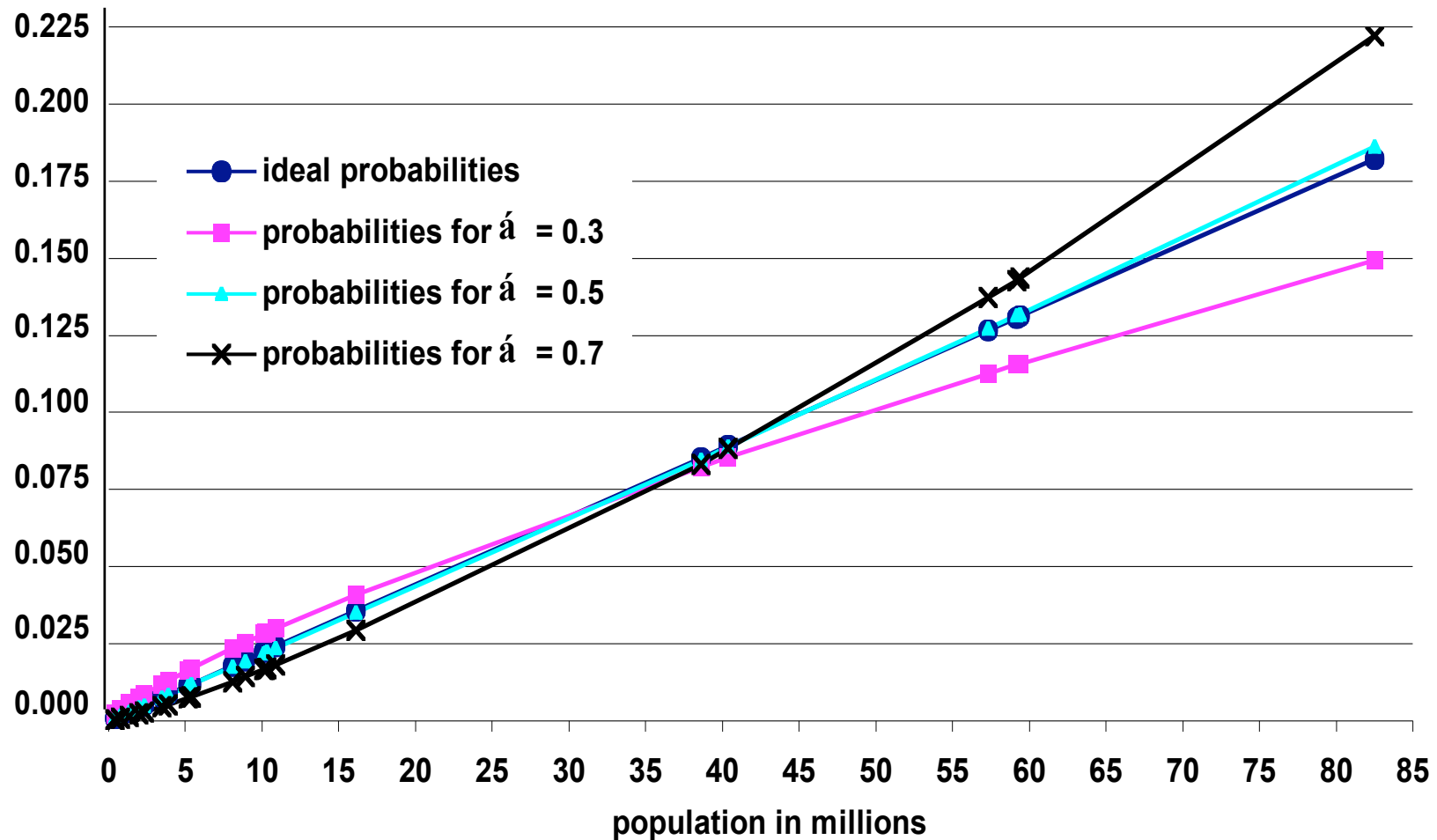
- Asymptotic approximation with only  $n_1$  varying and  $n_2 = \dots = n_m$  seems possible, but for general  $n_1, \dots, n_m$ ?

# Monte Carlo simulation

- Probability  $\pi_j := \Pr(j = (P))$  is the expected value of random variable  $H_j(\lambda_1, \dots, \lambda_m)$  which is 1 if  $j = (P)$  and 0 otherwise
- $H_j$ 's expected value can be approximated by the *empirical average* of many independent draws of  $H_j$
- Weight vectors are constructed from given population sizes by
$$w_j = n_j^\alpha$$
- For fixed weights  $(w_1, \dots, w_m)$  and populations  $(n_1, \dots, n_m)$ , we draw  $\lambda_1, \dots, \lambda_m$  from the beta distributions corresponding to i.i.d.  $U[0,1]$  voters in all constituencies and average  $H_1, \dots, H_m$  over 10 million draws
- We search for the  $\alpha$  which yields smallest cumulative (individual) quadratic deviation of  $\pi_j$  from the ideal egalitarian probability  $\pi_j^* = n_j / \sum n_k$  ( $j=1, \dots, m$ )

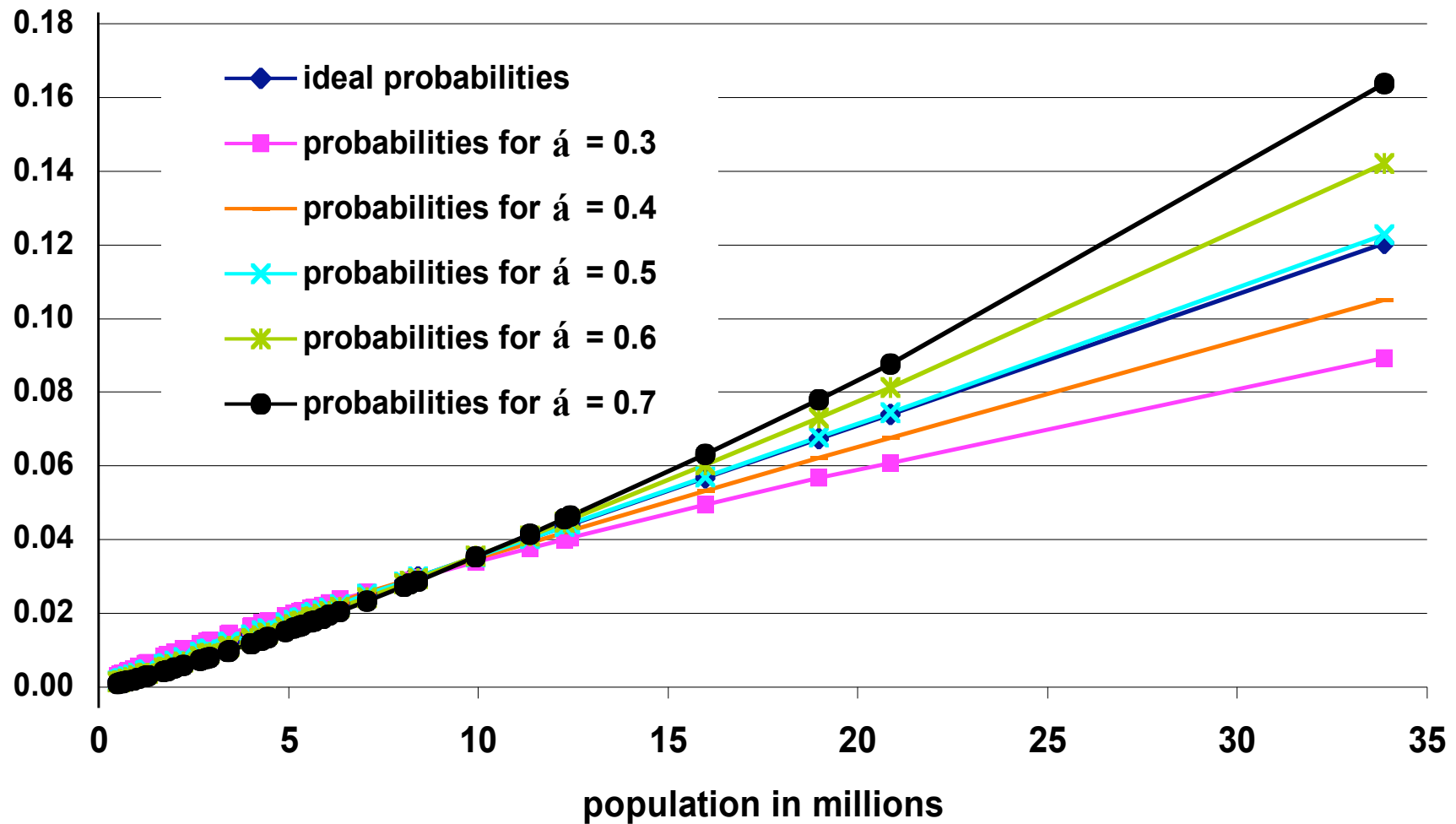
# EU Council of Ministers

- Using EU25 population data,  $\alpha^*=0.5$  with 50%-quota would give almost equal representation:



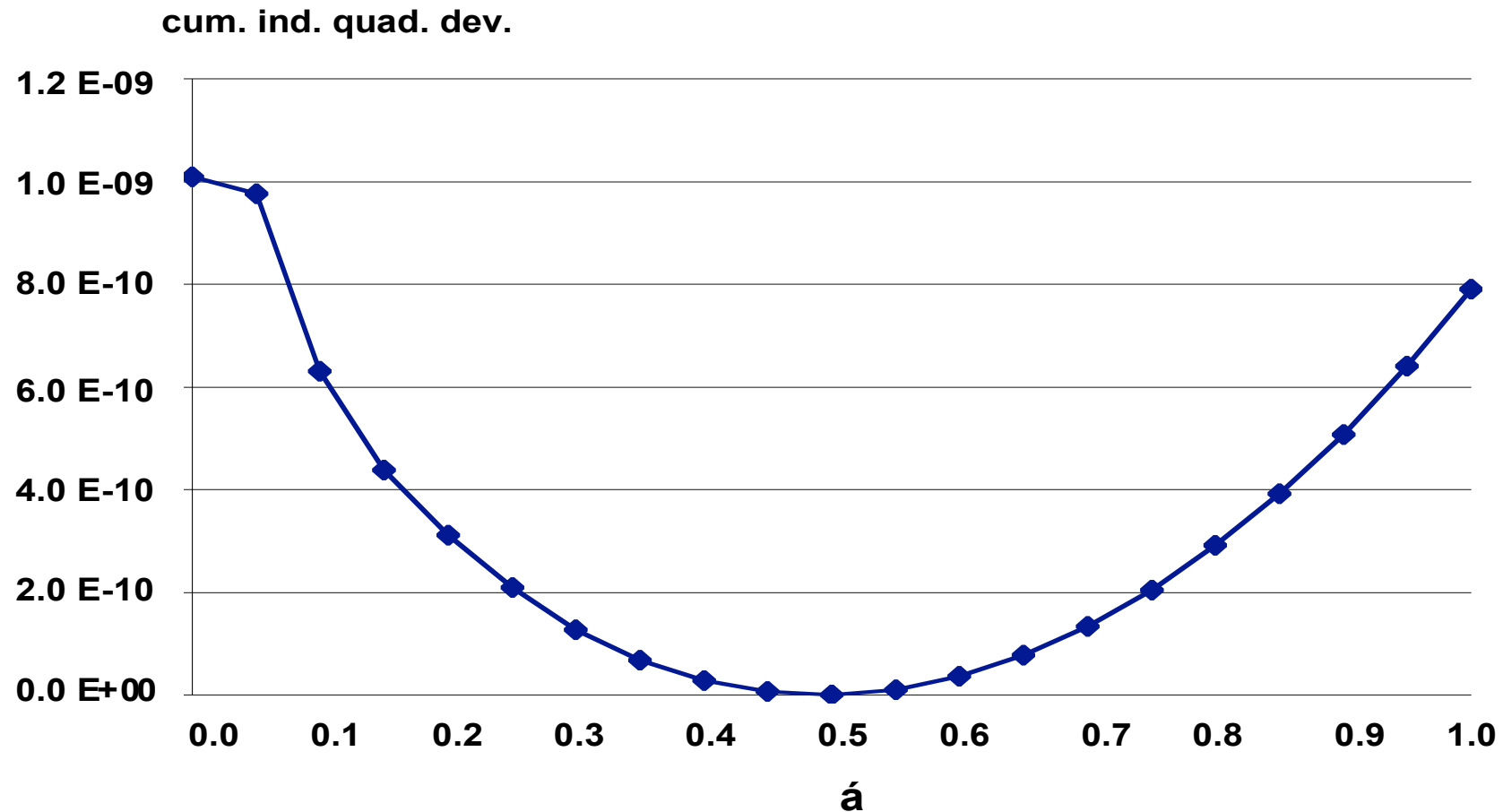
# US Electoral College

- Again,  $\alpha^*=0.5$  comes very close to equal representation:



# US Electoral College

- Cumulative individual quadratic deviation from equal representation in the US Electoral College:



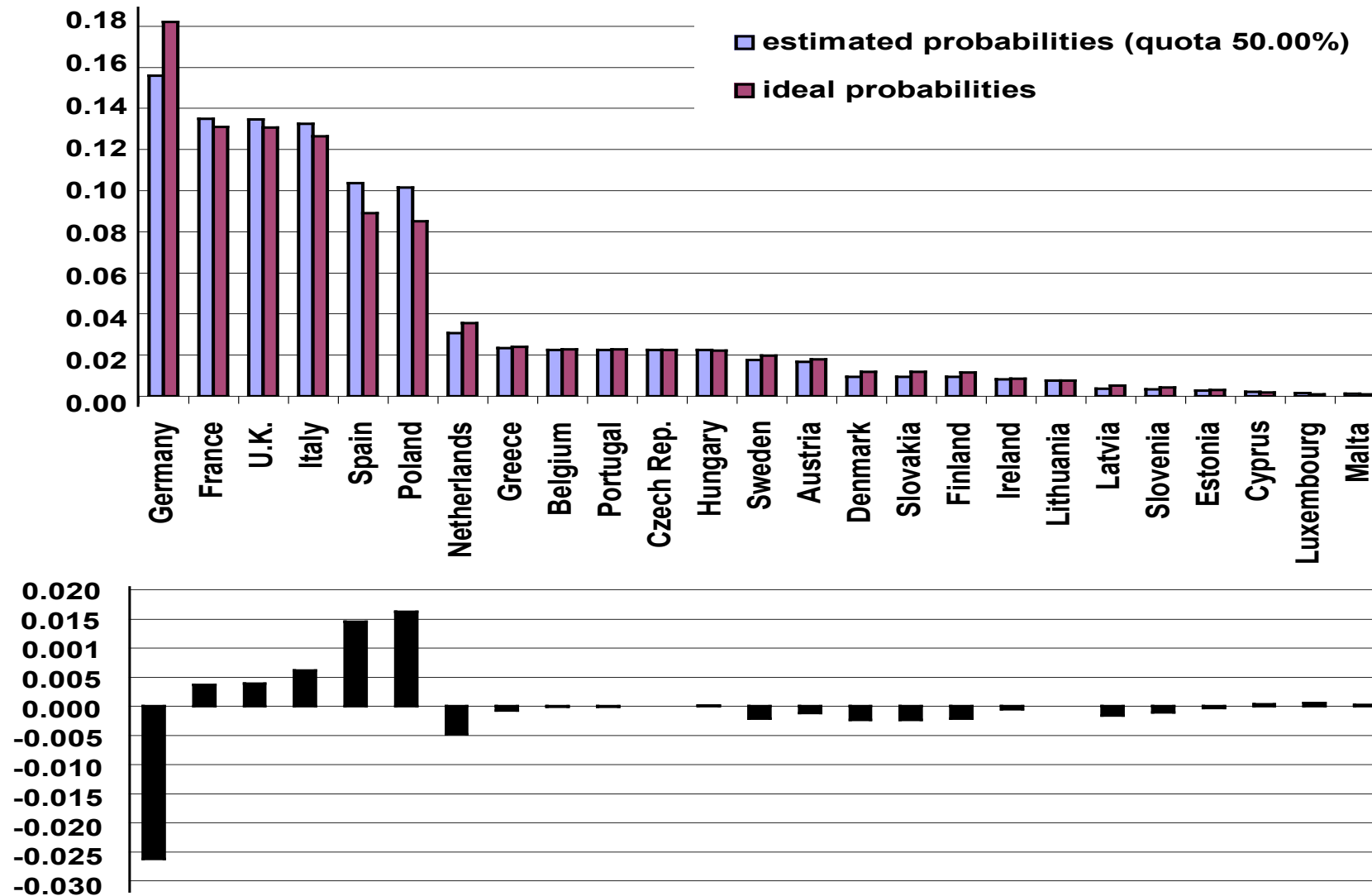
## Concluding remarks

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- While analytical proof of this looks out of reach, assigning *weights* proportional to *square root of population* provides a quite stable and satisfying answer to our question
- Thus, Penrose's square root rule is much more robust than suggested in the literature; unexpectedly, it extends from binary decisions to rich (one-dimensional convex) policy spaces, from simple games to spatial voting
- Future research:
  - A better reference point than voting weight
  - Effects of supermajority rule

# EU Council of Ministers

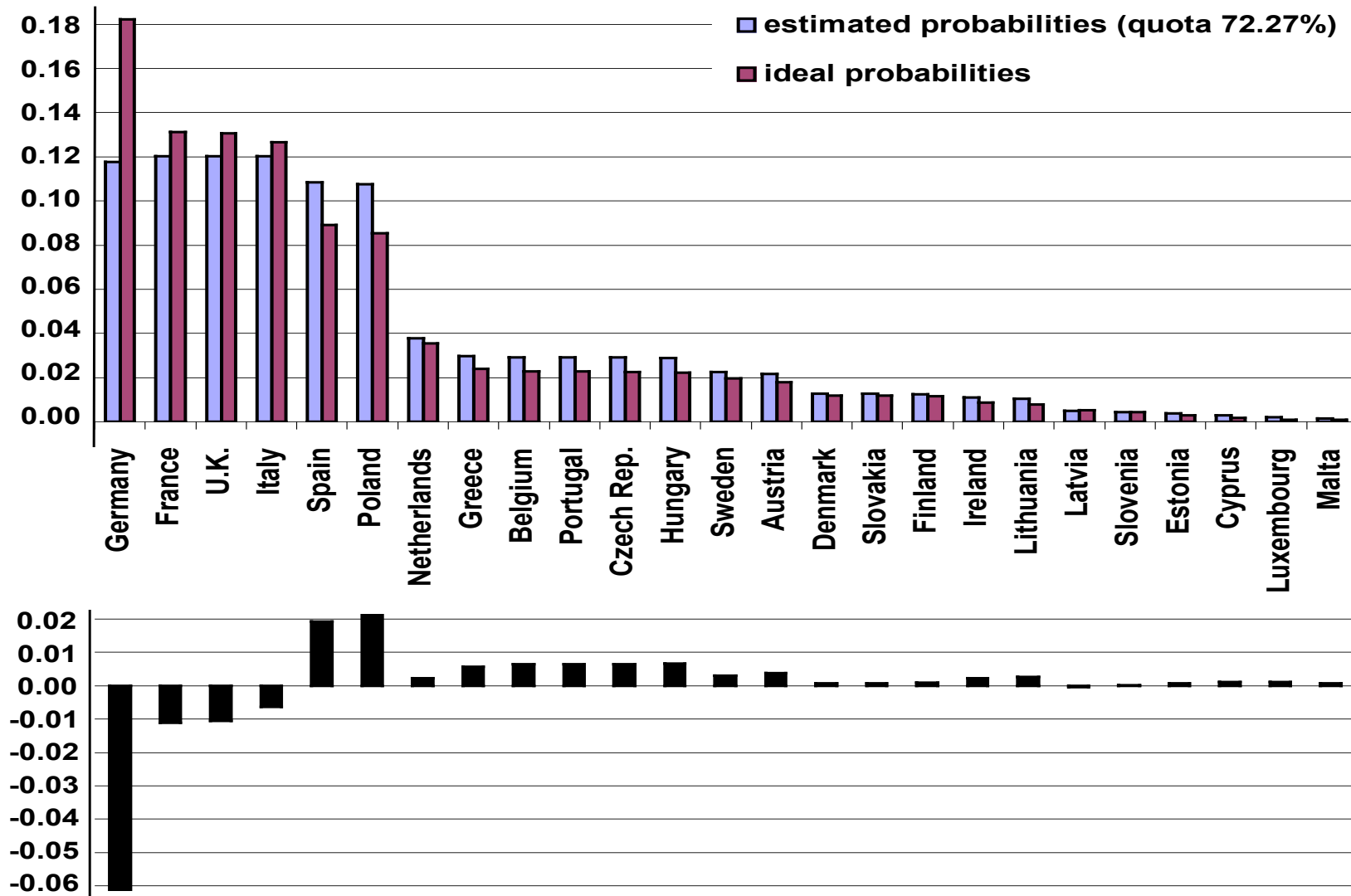
- Nice weights and quota of 50%:





# EU Council of Ministers

- Nice weights and quota of 72.2%:



## Results: uniformly distributed $n_j$

- We look, first, at  $m$  ranging from 10 to 50 with randomly generated constituency sizes  $n_1, \dots, n_m$  and, second, at two prominent real-world population configurations
- With i.i.d.  $n_j$  from  $U[0.5 \cdot 10^6, 99.5 \cdot 10^6]$ , optimal  $\alpha$  is:

# const	(I)	(II)	(III)	(IV)
10	<b>0.5</b> ( $1.22 \times 10^{-11}$ )	<b>0.6</b> ( $1.04 \times 10^{-11}$ )	<b>0.39</b> ( $2.20 \times 10^{-12}$ )	<b>0.00</b> ( $2.39 \times 10^{-11}$ )
15	<b>0.5</b> ( $1.43 \times 10^{-11}$ )	<b>0.5</b> ( $1.45 \times 10^{-13}$ )	<b>0.49</b> ( $2.79 \times 10^{-14}$ )	<b>0.48</b> ( $8.84 \times 10^{-14}$ )
20	<b>0.5</b> ( $4.80 \times 10^{-14}$ )	<b>0.5</b> ( $8.59 \times 10^{-14}$ )	<b>0.49</b> ( $5.66 \times 10^{-15}$ )	<b>0.49</b> ( $6.91 \times 10^{-15}$ )
25	<b>0.5</b> ( $9.25 \times 10^{-15}$ )	<b>0.5</b> ( $1.28 \times 10^{-14}$ )	<b>0.49</b> ( $5.37 \times 10^{-15}$ )	<b>0.49</b> ( $7.69 \times 10^{-15}$ )
30	<b>0.5</b> ( $1.11 \times 10^{-15}$ )	<b>0.5</b> ( $5.12 \times 10^{-15}$ )	<b>0.49</b> ( $7.36 \times 10^{-15}$ )	<b>0.49</b> ( $2.38 \times 10^{-15}$ )
40	<b>0.5</b> ( $3.38 \times 10^{-15}$ )	<b>0.5</b> ( $5.11 \times 10^{-15}$ )	<b>0.49</b> ( $3.69 \times 10^{-15}$ )	<b>0.49</b> ( $7.02 \times 10^{-15}$ )
50	<b>0.5</b> ( $3.06 \times 10^{-15}$ )	<b>0.5</b> ( $4.70 \times 10^{-15}$ )	<b>0.50</b> ( $3.10 \times 10^{-15}$ )	<b>0.50</b> ( $3.30 \times 10^{-15}$ )

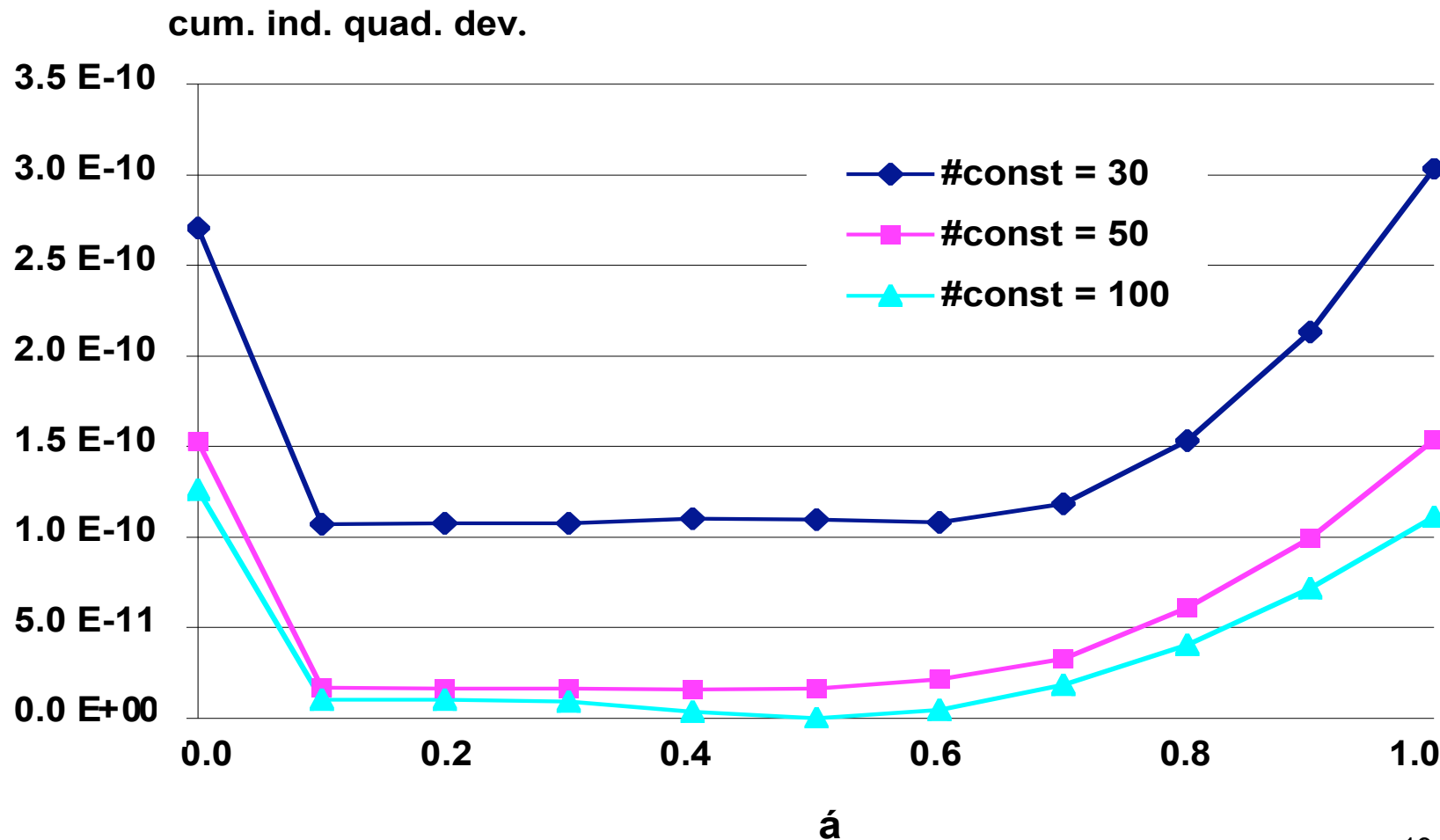
## Results: normally distributed $n_j$

- If constituencies are created for efficiency reasons, sizes possibly are distributed around some 'ideal size'
- With i.i.d.  $n_j$  from  $N(10^6; 200\,000)$ , optimal  $\alpha$  is:

# const	(I)	(II)	(III)	(IV)
10	<b>0.0</b> ( $1.22 \times 10^{-9}$ )	<b>0.0</b> ( $1.65 \times 10^{-9}$ )	<b>0.0</b> ( $9.21 \times 10^{-9}$ )	<b>0.0</b> ( $1.83 \times 10^{-9}$ )
20	<b>0.6</b> ( $2.19 \times 10^{-10}$ )	<b>0.0</b> ( $2.93 \times 10^{-10}$ )	<b>0.6</b> ( $2.82 \times 10^{-10}$ )	<b>0.0</b> ( $3.83 \times 10^{-10}$ )
30	<b>0.1</b> ( $1.07 \times 10^{-10}$ )	<b>0.2</b> ( $1.07 \times 10^{-10}$ )	<b>0.4</b> ( $6.94 \times 10^{-11}$ )	<b>0.5</b> ( $6.76 \times 10^{-11}$ )
40	<b>0.3</b> ( $1.72 \times 10^{-11}$ )	<b>0.4</b> ( $2.08 \times 10^{-11}$ )	<b>0.4</b> ( $2.32 \times 10^{-11}$ )	<b>0.5</b> ( $2.81 \times 10^{-13}$ )
50	<b>0.4</b> ( $1.60 \times 10^{-11}$ )	<b>0.2</b> ( $7.39 \times 10^{-12}$ )	<b>0.3</b> ( $3.56 \times 10^{-11}$ )	<b>0.3</b> ( $4.72 \times 10^{-11}$ )
100	<b>0.5</b> ( $1.01 \times 10^{-13}$ )	<b>0.5</b> ( $2.30 \times 10^{-12}$ )	<b>0.5</b> ( $1.99 \times 10^{-13}$ )	<b>0.5</b> ( $3.44 \times 10^{-13}$ )

## Results: normally distributed $n_j$

- For moderately many similar constituencies, weighted voting may allow only quite high (and flat) inequality of representation:



## Results: Pareto distributed $n_j$

- More realistically, with i.i.d.  $n_j$  from a Pareto distribution with skewness parameter  $k$ , optimal  $\alpha$  is:

$\kappa$	Number of constituencies					
	10	20	30	40	50	100
1.0	0.5 ( $1.32 \times 10^{-9}$ )	0.5 ( $6.99 \times 10^{-11}$ )	0.5 ( $1.32 \times 10^{-11}$ )	0.5 ( $1.87 \times 10^{-11}$ )	0.5 ( $1.31 \times 10^{-10}$ )	0.5 ( $3.79 \times 10^{-12}$ )
1.8	0.5 ( $3.25 \times 10^{-9}$ )	0.5 ( $4.78 \times 10^{-11}$ )	0.5 ( $2.41 \times 10^{-11}$ )	0.5 ( $2.25 \times 10^{-11}$ )	0.5 ( $1.86 \times 10^{-11}$ )	0.5 ( $1.04 \times 10^{-12}$ )
3.4	0.0 ( $3.72 \times 10^{-9}$ )	0.5 ( $5.64 \times 10^{-11}$ )	0.5 ( $2.41 \times 10^{-11}$ )	0.5 ( $3.27 \times 10^{-12}$ )	0.5 ( $2.67 \times 10^{-12}$ )	0.5 ( $8.88 \times 10^{-13}$ )
5.0	0.0 ( $1.08 \times 10^{-8}$ )	0.0 ( $3.61 \times 10^{-9}$ )	0.1 ( $1.03 \times 10^{-10}$ )	0.15 ( $2.85 \times 10^{-11}$ )	0.1 ( $1.91 \times 10^{-10}$ )	0.5 ( $7.54 \times 10^{-13}$ )

→ General finding:

As long as  $m \geq 15$ ,  $\alpha^*=0.5$  comes close to equal representation;  
it does best amongst all considered rules for large  $m$   
(and for small  $m$  if the electorate's partition is not too equal nor oceanic)